Mini-Course 07 Kalman Particle Filters

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Agenda

- State Estimation Problems & Kalman Filter
 - Henrique Massard
- Steady State Kalman Filter, Extended & Encented Kalman Filter
 - Cesar Cunha Pacheco
- Particle Filter
 - Wellington Bettencurte and Julio Dutra

- Introduction
 - State estimation problems, also designated as nonstationary inverse problems;
 - Available measured data is used together with prior knowledge about the physical phenomena and the measuring devices, in order to sequentially produce estimates of the desired dynamic variables;
 - This is accomplished in such a manner that the error is minimized **statistically**.

- Position of an aircraft
 - Time integration of velocity components since departure
 - Models aren't perfect
 - Measured with an GPS and altimeter
 - Measurement devices aren't perfect



 $a = \frac{dv}{dt} = a$

time ----

Model error

velocity

a = constant



 \bullet Consider the model for the evolution of the vector ${\bf x}$

$$\mathbf{x}_k = \mathbf{f}_k(\mathbf{x}_{k-1}, \mathbf{w}_{k-1})$$

• Measurements available and are related to \mathbf{x}_k as

$$\mathbf{y}_k = \mathbf{h}_k(\mathbf{x}_k, \mathbf{v}_k)$$

k=1,2,... denotes time instant

• Where:

 $\mathbf{x} \in R^{n_x}$ = state variables to be estimated

 $\mathbf{y} \in \boldsymbol{R}^{n_y}$ = measurements

 $\mathbf{w} \in R^{n_w}$ = state noise vector

 $\mathbf{v} \in R^{n_v}$ = measurement noise

 The state estimation problem aims at obtaining information about x_k based on the state evolution model and on the measurements given by the observation model.

Bayesian Framework

- The solution of the inverse problem within the Bayesian framework is recast in the form of statistical inference from the posterior probability density, which is the model for the conditional probability distribution of the unknown parameters given the measurements.
 - The measurement model incorporating the related uncertainties is called the likelihood, that is, the conditional probability of the measurements given the unknown parameters.
 - The model for the unknowns that reflects all the uncertainty of the parameters without the information conveyed by the measurements, is called the prior model.

Bayesian Framework

The formal mechanism to combine the new information (measurements)

with the previously available information (prior) is known as the **Bayes'**

theorem:

$$\pi_{posterior}(\mathbf{x}) = \pi(\mathbf{x}|\mathbf{y}) = \frac{\pi(\mathbf{x})\pi(\mathbf{y}|\mathbf{x})}{\pi(\mathbf{y})}$$

where $\pi_{posterior}(\mathbf{x})$ is the **posterior probability density**, $\pi(\mathbf{x})$ is the **prior density**, $\pi(\mathbf{y}|\mathbf{x})$ is the **likelihood function** and $\pi(\mathbf{y})$ is the **marginal probability density** of the measurements, which plays the role of a normalizing constant.

• State Evolution Model: $\mathbf{x}_k = \mathbf{f}_k (\mathbf{x}_{k-1}, \mathbf{w}_{k-1})$

• Observation Model:
$$\mathbf{y}_k = \mathbf{h}_k(\mathbf{x}_k, \mathbf{v}_k)$$

- The *evolution-observation model* is based on the following assumptions:
 - The sequence \mathbf{x}_k for k=1,2...., is a Markovian process, that is,

$$\pi(\mathbf{x}_{k}|\mathbf{x}_{0},\mathbf{x}_{1},\mathbf{x}_{2},\ldots,\mathbf{x}_{k-1}) = \pi(\mathbf{x}_{k}|\mathbf{x}_{k-1})$$

• The sequence \mathbf{y}_k for k=1,2,..., is a Markovian process with respect to the history of \mathbf{x}_k , that is,

$$\pi(\mathbf{y}_k|\mathbf{x}_0,\mathbf{x}_1,\mathbf{x}_2,\ldots,\mathbf{x}_k) = \pi(\mathbf{y}_k|\mathbf{x}_k)$$

• The sequence \mathbf{x}_k depends on the past observations only through its own history, that is,

$$\pi(\mathbf{x}_{k}|\mathbf{x}_{k-1},\mathbf{y}_{1:k-1}) = \pi(\mathbf{x}_{k}|\mathbf{x}_{k-1})$$

State Estimation Problem <u>State Evolution Model:</u> $\mathbf{x}_k = \mathbf{f}_k (\mathbf{x}_{k-1}, \mathbf{w}_{k-1})$

Observation Model:
$$\mathbf{y}_k = \mathbf{h}_k(\mathbf{x}_k, \mathbf{v}_k)$$

- time domain: used to classify the method of solution in the domain $\mathbf{y}_{1:kf} = \{\mathbf{y}_k, k=1,...,kf\}$
 - The **prediction problem**, concerned with the determination of $\pi(\mathbf{x}_k | \mathbf{y}_{1:k-1})$;
 - The filtering problem, concerned with the determination of $\pi(\mathbf{x}_k | \mathbf{y}_{1:k})$;
 - The fixed-lag smoothing problem, concerned with the determination of $\pi(\mathbf{x}_k | \mathbf{y}_{1:k+p})$, where $p \ge 1$ is the fixed lag;
 - The whole-domain smoothing problem, concerned with the determination of $\pi(\mathbf{x}_k|\mathbf{y}_{1:kf})$

- Filtering Problem
 - By assuming that $\pi(\mathbf{x}_0 | \mathbf{y}_0) = \pi(\mathbf{x}_0)$ is available, the posterior probability density $\pi(\mathbf{x}_k | \mathbf{y}_{1:k})$ is then obtained with Bayesian filters in two steps: **prediction and update**



- Evolution and observation models are linear.
- Noises in such models are additive and Gaussian, with known means and covariances.

<u>State Evolution Model:</u> $\mathbf{x}_k = \mathbf{F}_k \mathbf{x}_{k-1} + \mathbf{s}_{k-1} + \mathbf{w}_{k-1}$

<u>Observation Model:</u> $\mathbf{y}_k = \mathbf{H}_k \mathbf{x}_k + \mathbf{v}_k$

- The above set o equations leads to the optimum solution of the state variables;
- F and H are known matrices for the linear evolutions of the state x and of the observation y, respectively.
- Vector s is assumed to be a known input .
- Noises w and v have zero means and covariance matrices Q and R, respectively p(w)~N(0,Q) and p(v)~N(0,R)

- Derivation of the equations
 - The inverse problem solution can be obtained by

$$\widehat{\mathbf{x}}_n = \max_{\mathbf{x}_n} \left\{ \pi(\mathbf{x}_n | \mathbf{y}_{0:n}) \right\}$$

• Where the posterior pdf can be obtained by the following proportionality relation:

$$\pi(\mathbf{x}_n|\mathbf{y}_{0:n}) \propto \pi(\mathbf{y}_n|\mathbf{x}_n)\pi(\mathbf{x}_n|\mathbf{y}_{0:n-1})$$

• Reminding the evolution observation models:

$$\mathbf{x}_n = \mathbf{F}_n \mathbf{x}_{n-1} + \mathbf{w}_{n-1}$$
$$\mathbf{y}_n = \mathbf{H}_n \mathbf{x}_n + \mathbf{v}_n$$

Derivation of the equations

• Determine the **expectation** and the **covariance**

matrix for these distributions in order for the

posterior distribution to be fully characterized.

• For the likelihood the expectation is obtained by

$$\begin{split} \mathbb{E}\left[\mathbf{y}_{n} | \mathbf{x}_{n}\right] &= \mathbb{E}\left[\mathbf{H}_{n} \mathbf{x}_{n} + \mathbf{v}_{n}\right] & \text{using} \quad \mathbf{y}_{n} = \mathbf{H}_{n} \mathbf{x}_{n} + \mathbf{v}_{n} \\ &= \mathbf{H}_{n} \mathbf{x}_{n} \end{split}$$

• The covariance matrix is obtained by:

$$\operatorname{cov} \left[\mathbf{y}_n | \mathbf{x}_n \right] = \operatorname{cov} \left[\mathbf{v}_n \right]$$
$$= \mathbf{R}_n$$

- Derivation of the equations
 - Thus the pdf of the **likelihood** can be written as

$$\pi(\mathbf{y}_n|\mathbf{x}_n) = \frac{1}{(2\pi)^{N_y/2}|\mathbf{R}_n|^{1/2}} \exp\left[-\frac{1}{2}\left(\mathbf{y}_n - \mathbf{H}_n\mathbf{x}_n\right)^T \mathbf{R}_n^{-1}\left(\mathbf{y}_n - \mathbf{H}_n\mathbf{x}_n\right)\right]$$

- Derivation of the equations
 - For the **prior** pdf $\pi(\mathbf{x}_n | \mathbf{y}_{0:n-1})$, the expectation is obtained by:

$$\begin{split} \mathbb{E}\left[\mathbf{x}_{n} | \mathbf{y}_{0:n-1}\right] &= \mathbb{E}\left[\mathbf{F}_{n} \widehat{\mathbf{x}}_{n-1} + \mathbf{w}_{n-1}\right] \\ &= \mathbf{F}_{n} \widehat{\mathbf{x}}_{n-1} \\ &= \widehat{\mathbf{x}}_{n|n-1}^{\mathbf{\ddagger} \text{ Equivalent}} \end{split}$$

Both using

• And the covariance matrix is obtained by:

 $\mathbf{x}_n = \mathbf{F}_n \mathbf{x}_{n-1} + \mathbf{w}_{n-1}$

$$\operatorname{cov} [\mathbf{x}_{n} | \mathbf{y}_{0:n-1}] = \operatorname{cov} [\mathbf{x}_{n} - \widehat{\mathbf{x}}_{n|n-1}]$$

=
$$\operatorname{cov} [\mathbf{F}_{n} \mathbf{x}_{n-1} + \mathbf{w}_{n-1} - \mathbf{F}_{n} \widehat{\mathbf{x}}_{n-1}]$$

=
$$\operatorname{cov} [\mathbf{F}_{n} (\mathbf{x}_{n-1} - \widehat{\mathbf{x}}_{n-1}) + \mathbf{w}_{k-1}] \supset \operatorname{cov} [\mathbf{x}] = \mathbf{E} (\mathbf{x} \mathbf{x}^{\mathrm{T}})$$

=
$$\mathbf{F}_{n} \mathbf{P}_{n-1} \mathbf{F}_{n}^{\mathrm{T}} + \mathbf{Q}_{n}$$

=
$$\mathbf{P}_{n|n-1}^{\ddagger} \operatorname{Equivalent}$$

=
$$\mathbf{P}_{n|n-1}^{\ddagger}$$

- Derivation of the equations
 - Thus the pdf of the **priori** can be written as

$$\pi(\mathbf{x}_{n}|\mathbf{y}_{0:n-1}) = \frac{1}{(2\pi)^{N_{x}/2}|\mathbf{P}_{n|n-1}|^{1/2}} \exp\left[-\frac{1}{2}\left(\mathbf{x}_{n}-\widehat{\mathbf{x}}_{n|n-1}\right)^{T}\mathbf{P}_{n|n-1}^{-1}\left(\mathbf{x}_{n}-\widehat{\mathbf{x}}_{n|n-1}\right)\right]$$

- Derivation of the equations
 - The pdf of the posterior distribution is obtained combining both prior and likelihood pdf's:

$$\pi \left(\mathbf{x}_n | \mathbf{y}_{0:n} \right) \propto \exp \left\{ -\frac{1}{2} \left[\left(\mathbf{y}_n - \mathbf{H}_n \mathbf{x}_n \right)^T \mathbf{R}_n^{-1} \left(\mathbf{y}_n - \mathbf{H}_n \mathbf{x}_n \right) + \left(\mathbf{x}_n - \widehat{\mathbf{x}}_{n|n-1} \right)^T \mathbf{P}_{n|n-1}^{-1} \left(\mathbf{x}_n - \widehat{\mathbf{x}}_{n|n-1} \right) \right] \right\}$$

• The logarithm of the posterior pdf is taken and the derivative is calculated as

$$\frac{\partial}{\partial \mathbf{x}_n} \left[\ln \pi(\mathbf{x}_n | \mathbf{y}_{0:n}) \right] = \mathbf{H}_n^T \mathbf{R}_n^{-1} \left(\mathbf{y}_n - \mathbf{H}_n \mathbf{x}_n \right) - \mathbf{P}_{n|n-1}^{-1} \left(\mathbf{x}_n - \widehat{\mathbf{x}}_{n|n-1} \right)$$

- Derivation of the equations
 - Taking x_n as \hat{x}_n^{MAP} one obtains: [I]

$$\widehat{\mathbf{x}}_{n}^{\text{MAP}} = \left(\mathbf{H}_{n}^{T}\mathbf{R}_{n}^{-1}\mathbf{H}_{n} + \mathbf{P}_{n|n-1}^{-1}\right)^{-1} \left(\mathbf{P}_{n|n-1}^{-1}\widehat{\mathbf{x}}_{n|n-1} + \mathbf{H}_{n}^{T}\mathbf{R}_{n}^{-1}\mathbf{y}_{n}\right)$$

Π

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|IV| = |I||III|

• By using the following lemma for the first term in the right hand side (I),

$$\left(\mathbf{A} + \mathbf{C}\mathbf{B}\mathbf{C}^{T}\right)^{-1} = \mathbf{A}^{-1} - \mathbf{A}^{-1}\mathbf{C}\left(\mathbf{B}^{-1} + \mathbf{C}^{T}\mathbf{A}^{-1}\mathbf{C}\right)\mathbf{C}^{T}\mathbf{A}^{-1}$$

• Where

$$\mathbf{A} = \mathbf{P}_{n|n-1}^{-1}, \quad \mathbf{B} = \mathbf{R}^{-1}, \quad \mathbf{C} = \mathbf{H}_n^T$$

• One obtains,

$$\begin{pmatrix} \mathbf{H}_n^T \mathbf{R}_n^{-1} \mathbf{H}_n + \mathbf{P}_{n|n-1}^{-1} \end{pmatrix}^{-1} = \mathbf{P}_{n|n-1} - \\ - \mathbf{P}_{n|n-1} \mathbf{H}_n^T \left(\mathbf{H}_n \mathbf{P}_{n|n-1} \mathbf{H}_n^T + \mathbf{R}_n \right)^{-1} \mathbf{H}_n \mathbf{P}_{n|n-1}$$

- Derivation of the equations
 - A further simplification is obtained with the following lemma:

$$\left(\mathbf{P}^{-1} + \mathbf{B}^T \mathbf{R}^{-1} \mathbf{B}\right)^{-1} \mathbf{B}^T \mathbf{R}^{-1} = \mathbf{P} \mathbf{B}^T \left(\mathbf{B} \mathbf{P} \mathbf{B}^T + \mathbf{R}\right)^{-1} \qquad \mathbf{I} \mathbf{V} = \mathbf{I} \mathbf{I} \mathbf{I} \mathbf{I} \mathbf{I}$$

• Where

$$\mathbf{P}^{-1} = \mathbf{P}_{n|n-1}^{-1}, \quad \mathbf{B} = \mathbf{H}_n, \quad \mathbf{R} = \mathbf{R}_n$$

• Then

$$\left(\mathbf{H}_{n}^{T}\mathbf{R}_{n}^{-1}\mathbf{H}_{n} + \mathbf{P}_{n|n-1}^{-1}\right)^{-1}\mathbf{H}_{n}^{T}\mathbf{R}_{n}^{-1} = \mathbf{P}_{n|n-1}\mathbf{H}_{n}^{T}\left(\mathbf{H}_{n}\mathbf{P}_{n|n-1}\mathbf{H}_{n}^{T} + \mathbf{R}_{n}\right)^{-1}$$

- Derivation of the equations
 - Reminding: Derivative of $\ln(\pi(x|y))$ with respect to x

$$\widehat{\mathbf{x}}_{n}^{\text{MAP}} = \left(\mathbf{H}_{n}^{T}\mathbf{R}_{n}^{-1}\mathbf{H}_{n} + \mathbf{P}_{n|n-1}^{-1}\right)^{-1} \left(\mathbf{P}_{n|n-1}^{-1}\widehat{\mathbf{x}}_{n|n-1} + \mathbf{H}_{n}^{T}\mathbf{R}_{n}^{-1}\mathbf{y}_{n}\right)$$

• With the information of the previous lemmas:

$$\begin{pmatrix} \mathbf{H}_{n}^{T} \mathbf{R}_{n}^{-1} \mathbf{H}_{n} + \mathbf{P}_{n|n-1}^{-1} \end{pmatrix}^{-1} = \mathbf{P}_{n|n-1} - \\ - \mathbf{P}_{n|n-1} \mathbf{H}_{n}^{T} \left(\mathbf{H}_{n} \mathbf{P}_{n|n-1} \mathbf{H}_{n}^{T} + \mathbf{R}_{n} \right)^{-1} \mathbf{H}_{n} \mathbf{P}_{n|n-1} \\ \begin{pmatrix} \mathbf{H}_{n}^{T} \mathbf{R}_{n}^{-1} \mathbf{H}_{n} + \mathbf{P}_{n|n-1}^{-1} \end{pmatrix}^{-1} \mathbf{H}_{n}^{T} \mathbf{R}_{n}^{-1} = \mathbf{P}_{n|n-1} \mathbf{H}_{n}^{T} \left(\mathbf{H}_{n} \mathbf{P}_{n|n-1} \mathbf{H}_{n}^{T} + \mathbf{R}_{n} \right)^{-1} \\ \bullet \text{ Results in } \mathbf{\hat{X}}_{n}^{\mathbf{M} \mathbf{A} \mathbf{P}} = \mathbf{\hat{X}}_{n|n-1} + \mathbf{K}_{n} \left(\mathbf{y}_{n} - \mathbf{H}_{n} \mathbf{\hat{X}}_{n|n-1} \right) \\ \bullet \text{ Where } \mathbf{K}_{n} = \mathbf{P}_{n|n-1} \mathbf{H}_{n}^{T} \left(\mathbf{H}_{n} \mathbf{P}_{n|n-1} \mathbf{H}_{n}^{T} + \mathbf{R} \right)^{-1}$$

- Derivation of the equations
 - Covariance matrix derivation

$$\mathbf{P}_{n} = \operatorname{cov} \left[\mathbf{x}_{n} - \widehat{\mathbf{x}}_{n \mid n-1}^{\text{MAP}} \right]$$

= $\operatorname{cov} \left[\mathbf{x}_{n} - \widehat{\mathbf{x}}_{n \mid n-1} - \mathbf{K}_{n} \left(\mathbf{y}_{n} - \mathbf{H}_{n} \widehat{\mathbf{x}}_{n \mid n-1} \right) \right]$
= $\operatorname{cov} \left[\left(\mathbf{I} - \mathbf{K}_{n} \mathbf{H}_{n} \right) \left(\mathbf{x}_{n} - \widehat{\mathbf{x}}_{n \mid n-1} \right) + \mathbf{K}_{n} \mathbf{v}_{n} \right]$
= $\left(\mathbf{I} - \mathbf{K}_{n} \mathbf{H}_{n} \right) \mathbf{P}_{n \mid n-1} \left(\mathbf{I} - \mathbf{K}_{n} \mathbf{H}_{n} \right)^{T} + \mathbf{K}_{n} \mathbf{R}_{n} \mathbf{K}_{n}^{T}$
= $\left(\mathbf{I} - \mathbf{K}_{n} \mathbf{H}_{n} \right) \mathbf{P}_{n \mid n-1} \left(\mathbf{I} - \mathbf{K}_{n} \mathbf{H}_{n} \right)^{T} + \mathbf{K}_{n} \mathbf{R}_{n} \mathbf{K}_{n}^{T}$

• By using

$$\mathbf{K}_{n} = \mathbf{P}_{n|n-1} \mathbf{H}_{n}^{T} \left(\mathbf{H}_{n} \mathbf{P}_{n|n-1} \mathbf{H}_{n}^{T} + \mathbf{R} \right)^{-1}$$

• One obtain

$$\mathbf{P}_n = \left(\mathbf{I} - \mathbf{K}_n \mathbf{H}_n\right) \mathbf{P}_{n|n-1}$$

The Kalman Filter **Prediction:**

$$\hat{\mathbf{x}}_{k}^{-} = \mathbf{F}_{k} \hat{\mathbf{x}}_{k-1}^{+}$$
$$\mathbf{P}_{k}^{-} = \mathbf{F}_{k} \mathbf{P}_{k-1}^{+} \mathbf{F}_{k}^{T} + \mathbf{Q}_{k}$$



$$\mathbf{K}_{k} = \mathbf{P}_{k}^{-}\mathbf{H}_{k}^{T}\left(\mathbf{H}_{k}\mathbf{P}_{k}^{-}\mathbf{H}_{k}^{T} + \mathbf{R}_{k}\right)^{-1}$$
$$\hat{\mathbf{x}}_{k}^{+} = \hat{\mathbf{x}}_{k}^{-} + \mathbf{K}_{k}\left(\mathbf{y}_{k}^{-} - \mathbf{H}_{k}\hat{\mathbf{x}}_{k}^{-}\right)$$
$$\mathbf{P}_{k}^{+} = \left(\mathbf{I} - \mathbf{K}_{k}\mathbf{H}_{k}\right)\mathbf{P}_{k}^{-}$$
Measurement Innovation

 T_{∞} \boldsymbol{q} Heat flux h V Х

$$\frac{d\theta}{dt} + m\theta(t) = \frac{mq(t)}{h} \quad \text{For } t > 0$$

$$\theta(0) = \theta_o \quad \text{For } t = 0$$

where

$$\theta = T(t) - T_\infty \quad m = \frac{h}{\rho cL}$$

$$\theta_o = T_o - T_\infty$$

- Two illustrative cases are examined:
 - Heat flux q(t)=q_o constant and deterministically known;
 - Heat flux q(t)=q_of(t) with unknown time variation;
- Plate is made of **aluminum** (ρ = 2707 kgm⁻³, c = 896 Jkg⁻¹K⁻¹), with thickness L = 0.03 m, q₀ = 8000 W/m², T_∞=20 °C, h = 50 Wm⁻²K⁻¹ and T₀ = 50 °C.
- Measurement of the transient temperature of the slab are assumed available. These measurements contain additive, uncorrelated, Gaussian error, with zero mean and a constant standard deviation σ_z
- The errors in the state evolution model are also supposed to be additive, uncorrelated, Gaussian, with zero mean and a constant standard deviation σ_θ

- (i) Heat Flux $q(t)=q_o$ constant and deterministically known
 - The analytical solution for this problem is given by:

$$\theta(t) = \theta_o e^{-mt} + \frac{q_o}{h} \left(1 - e^{-mt} \right)$$

• The only state variable in this case is the temperature $\theta(t_k) = \theta_k$ since the applied heat flux q_o is constant and deterministically known, as the other parameter appearing in the formulation. By using a forward finite differences approximation for the time derivative in equation:

$$\frac{d\theta}{dt} + m\theta(t) = \frac{mq(t)}{h}$$

We obtain:

$$\theta_{k} = \left(1 - m\Delta t\right)\theta_{k-1} + \frac{mq_{o}}{h}\Delta t$$

• Therefore, the state and observation models given by:

$$\mathbf{x}_{k} = \mathbf{F}_{k} \mathbf{x}_{k-1} + \mathbf{s}_{k-1} + \mathbf{w}_{k-1}$$

$$\mathbf{z}_k = \mathbf{H}_k \mathbf{x}_k + \mathbf{n}_k$$

are obtained with

$$\mathbf{x}_{k} = \begin{bmatrix} \theta_{k} \end{bmatrix} \qquad \mathbf{s}_{k} = \begin{bmatrix} m \frac{q_{o}}{h} \Delta t \end{bmatrix}$$
$$\mathbf{F}_{k} = \begin{bmatrix} (1 - m\Delta t) \end{bmatrix} \qquad \mathbf{H}_{k} = \begin{bmatrix} 1 \end{bmatrix}$$

$$\mathbf{Q}_k = \left[\boldsymbol{\sigma}_{\theta}^2 \right]$$

$$\mathbf{R}_{k} = \left[\sigma_{z}^{2} \right]$$

- (ii) Heat Flux $q(t) = q_0 f(t)$ with unknown time variation,
 - The analytical solution for this problem is given by:

$$\theta(t) = \left\{ \theta_o + \frac{mq_o}{h} \int_{t'=0}^t e^{mt'} f(t') dt' \right\}$$

• In this case, the state variables are given by the temperatures $\theta(t_k) = \theta_k$ and the function that gives the time variation of the applied heat flux, that is $f(t_k) = f_k$. As in the case examined previously, the applied heat flux \mathbf{q}_o is constant and deterministically known, as the other parameters appearing in the formulation. By using a forward finite-difference approximation for the time derivative we obtain the equation for the evolution of the state variable $\theta(t_k) = \theta_k$:

$$\theta_k = (1 - m\Delta t)\theta_{k-1} + \left(\frac{mq_o}{h}\Delta t\right)f_{k-1}$$

• A random walk model is used for the state variable $f(t_k)=f_{k_1}$, which is given in the form

$$f_k = f_{k-1} + \mathcal{E}_{k-1}$$

- Where $\epsilon_{k\text{-1}}$ is Gaussian with zero mean and constant standard deviation $\sigma_{\!rw}$

- (ii) Heat Flux $q(t) = q_o f(t)$ with unknown time variation
 - Therefore, the state and observation models given by:

$$\mathbf{x}_k = \mathbf{F}_k \mathbf{x}_{k-1} + \mathbf{s}_{k-1} + \mathbf{w}_{k-1}$$

$$\mathbf{z}_k = \mathbf{H}_k \mathbf{x}_k + \mathbf{n}_k$$

are obtained with



 $\mathbf{F}_{k} = \begin{bmatrix} (1 - m\Delta t) & \frac{mq_{o}}{h}\Delta t \\ 0 & 1 \end{bmatrix} \quad \mathbf{X}_{k} = \begin{bmatrix} \theta_{k} \\ f_{k} \end{bmatrix}$ $\mathbf{H}_{k} = \begin{bmatrix} 1 & 0 \end{bmatrix}$

$$\mathbf{Q}_{k} = \begin{bmatrix} \sigma_{\theta}^{2} & 0 \\ 0 & \sigma_{rw}^{2} \end{bmatrix}$$

$$\mathbf{R}_{k} = \left[\sigma_{z}^{2}\right]$$

• Estimation of position-dependent transient heat source

$$C\frac{\partial \overline{T}}{\partial t} = \frac{\partial}{\partial x} \left(k \frac{\partial \overline{T}}{\partial x} \right) + \frac{\partial}{\partial y} \left(k \frac{\partial \overline{T}}{\partial y} \right) - \frac{h}{e} \left(\overline{T} - T_{\infty} \right) + \frac{q(x, y, t)}{e} \quad \text{in } 0 < x < a, 0 < y < b, \text{ for } t > 0$$

$$\frac{\partial \overline{T}}{\partial y} = 0 \qquad \text{at } x = 0 \text{ and } x = a, \text{ for } t > 0$$
$$\frac{\partial \overline{T}}{\partial x} = 0 \qquad \text{at } y = 0 \text{ and } y = b, \text{ for } t > 0$$

 $\overline{T} = T_0$ for t = 0, in 0 < x < a and 0 < y < b



• Schematics of the experiment



Nodal approach



$$\frac{\partial T}{\partial t} = \alpha(x, y) \nabla^2 T - H(x, y) \left(T - T_{\infty} \right) + G(x, y, t)$$

$$Y_{i,j}^{n+1} = \Delta t \left(\frac{T_{i-1,j}^n - 2T_{i,j}^n + T_{i+1,j}^n}{(\Delta x)^2} + \frac{T_{i,j-1}^n - 2T_{i,j}^n + T_{i,j+1}^n}{(\Delta y)^2} \right) \alpha_{i,j} - \Delta t (T_{i,j}^n - T_{\infty}) H_{i,j} + \Delta t G_{i,j}$$



• Therefore, the state and observation models given by:

$$\mathbf{x}_{k} = \mathbf{F}_{k} \mathbf{x}_{k-1} + \mathbf{s}_{k-1} + \mathbf{w}_{k-1}$$

$$\mathbf{z}_k = \mathbf{H}_k \mathbf{x}_k + \mathbf{n}_k$$

are obtained with

$$\mathbf{x}_{k} = \begin{bmatrix} \mathbf{T}_{k} \\ \mathbf{P}_{k} \end{bmatrix}$$

 $\mathbf{F}^{k} = \begin{bmatrix} \mathbf{I} & \mathbf{J}^{k} \\ \mathbf{0} & \mathbf{I} \end{bmatrix}$

 $\mathbf{H}^{k} = [\mathbf{I} \ \mathbf{0}]$

$$\mathbf{Q}_k = \begin{bmatrix} \mathbf{Q}_T^k & \mathbf{0} \\ \mathbf{0} & \mathbf{Q}_P^k \end{bmatrix}$$
Example II: Experimental result

 ΔT (°C) in time step = 30 2.8 2.6 2.4 П 2.2 1.8 Sinusoidal pulse



Example II: Experimental result - Sine



Example II: Experimental result - Square



Example II: Experimental result - Sine



Example II: Experimental result - Square



Example II: Experimental result





The Kalman Filter

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The Kalman Filter

• Thank you!

Fact: Kalman filter is restrict to linear and Gaussian problems.

One investigation:

- System matrices are allowed to change with time.
- What if they don't?

Second investigation:

- What if the hypothesis above do not hold?
- Can we solve **nonlinear and/or non Gaussian** problems?

• What if instead of

$$\mathbf{x}_{k} = \mathbf{F}_{k}\mathbf{x}_{k-1} + \mathbf{u}_{k-1} + \mathbf{w}_{k-1}$$

$$\mathbf{x}_{k} = \mathbf{F}\mathbf{x}_{k-1} + \mathbf{u}_{k-1} + \mathbf{w}_{k-1}$$

$$\mathbf{y}_k = \mathbf{H}_k \mathbf{x}_k + \mathbf{n}_k$$

$$\mathbf{y}_k = \mathbf{H}\mathbf{x}_k + \mathbf{n}_k$$

with the noise given by

$$\mathbf{w}_k \sim N(\mathbf{0}, \mathbf{Q}) \qquad \mathbf{n}_k \sim N(\mathbf{0}, \mathbf{R})$$

In other words

$$\mathbf{F}_k = \mathbf{F}, \quad \mathbf{H}_k = \mathbf{H}, \quad \mathbf{Q}_k = \mathbf{Q}, \quad \mathbf{R}_k = \mathbf{R}$$

As a practical result, **K** and **P** matrices behave asymptotically. After a number of steps:

$$\mathbf{K}_{k} \simeq \mathbf{K}_{\infty}, \quad \mathbf{P}_{k}^{-} \simeq \mathbf{P}_{k}^{+} \simeq \mathbf{P}_{\infty}$$

Can we use this to our advantage?

Steady-State Kalman Filter

Applying this result to Kalman filter equations yields:

$$\mathbf{P}_{\infty} = \mathbf{F}\mathbf{P}_{\infty}\mathbf{F}^{T} - \mathbf{F}\mathbf{P}_{\infty}\mathbf{H}^{T} \left(\mathbf{H}\mathbf{P}_{\infty}\mathbf{H}^{T} + \mathbf{R}\right)^{-1}\mathbf{H}\mathbf{P}_{\infty}\mathbf{F}^{T} + \mathbf{Q}$$

$$\mathbf{K}_{\infty} = \mathbf{P}_{\infty} \mathbf{H}^{T} \left(\mathbf{H} \mathbf{P}_{\infty} \mathbf{H}^{T} + \mathbf{R} \right)^{-1}$$

$$\hat{\mathbf{x}}_{n}^{+} = (\mathbf{I} - \mathbf{K}_{\infty}\mathbf{H})\mathbf{F}\hat{\mathbf{x}}_{n-1}^{+} + \mathbf{K}_{\infty}\mathbf{y}_{n}$$

These equations are referred to as the Steady-State Kalman Filter (SSKF).

Steady-State Kalman Filter

What are the properties of this method?

- Equations 1 and 2: offline;
- Equation 3: *online*; Thus:

$$\mathbf{P}_{\infty} = \mathbf{F} \mathbf{P}_{\infty} \mathbf{F}^{T} - \mathbf{F} \mathbf{P}_{\infty} \mathbf{H}^{T} \left(\mathbf{H} \mathbf{P}_{\infty} \mathbf{H}^{T} + \mathbf{R} \right)^{-1} \mathbf{H} \mathbf{P}_{\infty} \mathbf{F}^{T} + \mathbf{Q}$$

n

- **P** and **K** are calculated *offline*;
- **x** and **y** appear in one equation;
- No online matrix inversion;
- O(n²), instead of O(n³);

$$\mathbf{K}_{\infty} = \mathbf{P}_{\infty} \mathbf{H}^{T} \left(\mathbf{H} \mathbf{P}_{\infty} \mathbf{H}^{T} + \mathbf{R} \right)^{-1}$$
$$\hat{\mathbf{X}}_{n}^{+} = \left(\mathbf{I} - \mathbf{K}_{\infty} \mathbf{H} \right) \mathbf{F} \hat{\mathbf{X}}_{n-1}^{+} + \mathbf{K}_{\infty} \mathbf{y}_{n}$$

Example III – Introduction

Heat Flux Identification problem:

- Heating of a flat square plate;
- 3D Nonlinear Heat Conduction;
- High magnitude heat flux;
- Measurements taken at opposite side;
- (a,b,c)=(120,120,30) mm



3D Nonlinear Transient Inverse Heat Conduction Problem: How to estimate the heat flux at real time?

Complete Mathematical Model:

• Governing Equation (t > 0):

$$C(T_c)\frac{\partial T_c}{\partial t} = \nabla \cdot \left[k_c(T_c)\nabla T_c\right],$$



• Initial Condition (t = 0):

 $T(x, y, z, t) = T_0,$

Thermal Properties $C(T) = 1324.75T + 3557900 \text{ [J/m}^3\text{]}$ $k(T) = 12.45 + 0.014T + 2.5171 \times 10^{-6} \text{ T}^2 \text{ [W/mK]}$

Complete Mathematical Model:

• Boundary Conditions

in
$$x = 0$$
 and $x = a$: $\frac{\partial T_c}{\partial x} = 0$
in $y = 0$ and $y = b$: $\frac{\partial T_c}{\partial y} = 0$

$$\begin{array}{c} & y \\ & & & & \\ & & & \\ & & & \\ & & & \\ & & & & \\ & & & \\ & & & & & \\ & & & & & \\ & & & & & \\ & & & & & \\ & & & & & \\ & & & & & \\ & &$$

in
$$z = 0$$
: $\frac{\partial T_c}{\partial z} = 0$ in $z = c$: $k(T_c)\frac{\partial T_c}{\partial z} = q(x, y, t)$

Complete Mathematical Model:

- Nonlinear: unable to apply Kalman filter for this case;
- 3D inverse problem might lead to high computational effort;

Proposal: Use a reduced model:

- Linearize it;
- Use mean temperature at the z-direction instead of actual temperature;
- Approximate temperature gradients across the thickness.

$$\left|\overline{T}(x, y, t) = \frac{1}{c} \int_{z=0}^{c} T(x, y, z, t) dz\right|$$

Reduced Mathematical Model:

• Governing Equation (t > 0):

$$C^* \frac{\partial \overline{T}}{\partial t} = k^* \frac{\partial^2 \overline{T}}{\partial x^2} + k^* \frac{\partial^2 \overline{T}}{\partial y^2} + \frac{q(x, y, t)}{c},$$

• Initial Condition (t = 0):

$$\overline{T}(x, y, t) = T_0,$$

Thermal Properties

$$C^* = C(T^*)$$

 $w/T^* = 600 \text{ K}$
 $k^* = k(T^*)$

Reduced Mathematical Model:

• Boundary Conditions

in
$$x = 0$$
 and $x = a$: $\frac{\partial T}{\partial x} = 0$
in $y = 0$ and $y = b$: $\frac{\partial \overline{T}}{\partial y} = 0$



Improved Lumped Method

• Temperature at z=0:

$$T(x, y, 0, t) \simeq \overline{T}(x, y, t) - \frac{c}{6k^*}q(x, y, t)$$

State vector:

- Mean temperature;
- Heat flux;
- Values throughout the mesh.

Observation vector:

- Temperature at z=0;
- Values throughout the mesh.

Noise Covariance Matrices:

• Assumed uncorrelated.





Synthetic Measurements:

- Solution of forward complete problem;
- Achieved grid/time-step independence;
- Exact (Reference) heat flux;
- <u>Total time: 2.0 s</u> (200 measurements); Inverse Problem:
- Reduced model;
- 24 x 24 grid with 0.02 s time step;
- No inverse crime;



OBS: $\boldsymbol{\omega} \sim N(\mathbf{0}, \mathbf{I})$

Exact (Reference) Heat Flux:

$$q(x, y, t) = \begin{cases} q_1, & x_{1,1} \le x \le x_{1,2}, \\ y_{1,1} \le y \le y_{1,2}, \\ q_2 & x_{2,1} \le x \le x_{2,2}, \\ q_2 & y_{2,1} \le y \le y_{2,2}, \\ 0, & \text{otherwise} \end{cases} t \ge t_0 \qquad \begin{bmatrix} q_1 & 1E7 \text{ W/m}^2 \\ q_2 & 5E6 \text{ W/m}^2 \\ t_1 & 0.4 \text{ s} \\ t_2 & 0.6 \text{ s} \end{bmatrix}$$

90 mm

100 mm

90 mm

100 mm



Example III – Results

Hot Spot #1: Temperature and heat flux estimates



Example III – Results

Hot Spot #1: Temperature estimates and residuals



Hot Spot #2: Temperature and heat flux estimates



Example III – Results

Hot Spot #2: Temperature and heat flux estimates



Example III – Results

Computational time:

- Physical time: 2.0 s;
- Kalman filter: 300 s;
- <u>SSKF: 0.9 s;</u>

Only SSKF allows for real-time estimation;

- Majority of computational effort done at pre-processing;
- Recursive estimation: 2 matrix-vector multiplications.

• What if instead of

$$\mathbf{x}_k = \mathbf{F}_k \mathbf{x}_{k-1} + \mathbf{u}_{k-1} + \mathbf{w}_{k-1}$$

$$\mathbf{y}_k = \mathbf{H}_k \mathbf{x}_k + \mathbf{n}_k$$

we had a nonlinear non Gaussian model:

$$\mathbf{x}_{k} = \mathbf{f}_{k} \left(\mathbf{x}_{k-1}, \mathbf{u}_{k-1}, \mathbf{w}_{k-1} \right)$$

$$\mathbf{y}_k = \mathbf{h}(\mathbf{x}_k, \mathbf{n}_k)$$

How does one perform sequential estimation with such models?

The Extended KF

• Idea: Linearize the model and apply classical KF equations:

$$\mathbf{x}_{k} = \mathbf{f}_{k} \left(\hat{\mathbf{x}}_{k-1}^{+}, \mathbf{u}_{k-1}^{-}, \mathbf{w}_{n-1}^{-} \right) + \frac{\partial \mathbf{f}_{k}}{\partial \mathbf{x}} \Big|_{\hat{\mathbf{x}}_{k-1}^{+}} \left(\mathbf{x}_{k-1}^{-} - \hat{\mathbf{x}}_{k-1}^{+} \right) + \frac{\partial \mathbf{f}_{k}}{\partial \mathbf{w}} \Big|_{\hat{\mathbf{x}}_{k-1}^{+}} \mathbf{w}_{k-1}^{-}$$

$$\mathbf{y}_{k} = \mathbf{h}_{k} \left(\hat{\mathbf{x}}_{k}^{-}, \mathbf{n}_{n} \right) + \frac{\partial \mathbf{h}_{k}}{\partial \mathbf{x}} \Big|_{\hat{\mathbf{x}}_{k}^{-}} \left(\mathbf{x}_{k} - \hat{\mathbf{x}}_{k}^{-} \right) + \frac{\partial \mathbf{h}_{k}}{\partial \mathbf{n}} \Big|_{\hat{\mathbf{x}}_{k}^{-}} \mathbf{n}_{k}$$

The Extended KF – Equations

Extended KF (EKF): Linearization + Kalman Filter:

• Linearization of both evolution and observation models.

Prediction:

• Linearization:

$$\mathbf{F}_{k-1} = \frac{\partial \mathbf{f}_k}{\partial \mathbf{x}}\Big|_{\hat{\mathbf{x}}_{k-1}^+} \text{ and } \mathbf{L}_{k-1} = \frac{\partial \mathbf{f}_k}{\partial \mathbf{w}}\Big|_{\hat{\mathbf{x}}_{k-1}^+}$$

• Prior Mean and Covariance

 $\hat{\mathbf{x}}_n^- = \mathbf{f}_k \left(\hat{\mathbf{x}}_{k-1}^+, \mathbf{u}_{k-1}^-, \mathbf{0} \right)$ and $\mathbf{P}_n^- = \mathbf{F}_{k-1}^- \mathbf{P}_{k-1}^+ \mathbf{F}_{k-1}^T + \mathbf{L}_{k-1}^- \mathbf{Q}_{k-1}^- \mathbf{L}_{k-1}^T$

The Extended KF – Equations

Update:

Linearization

$$\left| \mathbf{H}_{k} = \frac{\partial \mathbf{h}_{k}}{\partial \mathbf{x}} \right|_{\hat{\mathbf{x}}_{k}^{-}} \text{ and } \mathbf{M}_{k} = \frac{\partial \mathbf{h}_{k}}{\partial \mathbf{n}} \right|_{\hat{\mathbf{x}}_{k}^{-}}$$

Calculation of Intermediate Covariance Matrices

$$\mathbf{K}_{k} = \mathbf{P}_{k}^{-}\mathbf{H}_{k}^{T}\left(\mathbf{H}_{k}\mathbf{P}_{k}^{-}\mathbf{H}_{k}^{T} + \mathbf{M}_{k}\mathbf{R}_{k}\mathbf{M}_{k}^{T}\right)^{-1}$$
$$\hat{\mathbf{x}}_{k}^{+} = \hat{\mathbf{x}}_{k}^{-} + \mathbf{K}_{k}\left[\mathbf{y}_{k} - \mathbf{h}_{k}\left(\hat{\mathbf{x}}_{k}^{-}, \mathbf{0}\right)\right] \quad \text{and} \quad \mathbf{P}_{k}^{+} = \left(\mathbf{I} - \mathbf{K}_{k}\mathbf{H}_{k}\right)\mathbf{P}_{k}^{-1}$$



Free falling of a particle:

State variables:

- Altitude (x₁);
- Velocity (x₂);
- Ballistic Coefficient (x₃);
- Observation variable:
- Observed distance:

$$\sqrt{M^2 + \left[x_1(t) - a\right]^2}$$



Example IV



Example IV

Free falling of a particle: Input data:

$$\rho_0 = 2 \text{ lbs}^2/\text{ft}^4;$$

 $g = 32.2 \text{ ft/s}^2;$
 $k = 2 \times 10^4 \text{ ft};$
 $M = 10^5 \text{ ft};$
 $a = 10^5 \text{ ft}.$

Noise model:

$$E\left[w_i^2(t)\right] = 0, \quad i = 1, 2, 3$$
$$E\left[v_k^2\right] = 10^4 \text{ ft}^2$$

Initial State:

$$\hat{\mathbf{x}}_{0}^{+} = \mathbf{x}_{0} = \begin{bmatrix} 3 \times 10^{5} \\ -2 \times 10^{4} \\ 10^{-3} \end{bmatrix} \quad \mathbf{P}_{0}^{+} = \text{diag} \begin{bmatrix} 10^{6} \\ 4 \times 10^{6} \\ 10 \end{bmatrix}$$

Example IV

Free falling of a particle:

Reference values:





Estimation error:
EKF properties:

- Sequential nonlinear estimation algorithm of choice;
- Non-intrusive;
- Up for parallelization;
- Hard to accelerate theoretically (there is no "SSEKF").

What if nonlinearities are not sufficiently captured?

- 1st order approximations might be insufficient;
- EKF might lead to unreliable estimates.

Example: Polar to Rectangular mapping:

$$\mathbf{x} = \begin{bmatrix} r \\ \theta \end{bmatrix} \text{ and } \mathbf{y} = \begin{bmatrix} r \cos \theta \\ r \sin \theta \end{bmatrix}$$

Goal: Find **y** statistics, given

- Nonlinear mapping;
- **x** is uncorrelated;
- Symmetric pdfs.

$$\overline{\mathbf{x}} = \begin{bmatrix} \overline{r} \\ \overline{\theta} \end{bmatrix} = \begin{bmatrix} 1 \\ \frac{\pi}{2} \end{bmatrix} \text{ and } \mathbf{y} = \mathbf{h}(\mathbf{x})$$

• 1st order expansion around mean of **x**:

$$\overline{\mathbf{y}} = \mathbf{E}\left[\mathbf{h}(\mathbf{x})\right] \simeq \mathbf{E}\left[\mathbf{h}(\overline{\mathbf{x}}) + \frac{\partial \mathbf{h}}{\partial \mathbf{x}}\Big|_{\overline{\mathbf{x}}} \left(\mathbf{x} - \overline{\mathbf{x}}\right)\right]$$

• Thus

$$\overline{\mathbf{y}} \simeq \mathbf{h}(\overline{\mathbf{x}}) + \frac{\partial \mathbf{h}}{\partial \mathbf{x}}\Big|_{\overline{\mathbf{x}}} \mathbf{E}[\mathbf{x} - \overline{\mathbf{x}}] \quad \therefore \quad \overline{\mathbf{y}} \simeq \mathbf{h}(\overline{\mathbf{x}}) = \begin{bmatrix} 0\\1 \end{bmatrix}$$

1st component:

 $\overline{y}_{1} = \mathbb{E}[r\cos\theta]$ $= \mathbb{E}[(\overline{r} + \tilde{r})\cos(\overline{\theta} + \tilde{\theta})]$ $= \mathbb{E}[(\overline{r} + \tilde{r})(\cos\overline{\theta}\cos\overline{\theta} - \sin\overline{\theta}\sin\overline{\theta})]$ $= \overline{r}\cos\overline{\theta}$ = 0

Assuming:

$$\mathbf{x} = \begin{bmatrix} r \\ \theta \end{bmatrix} = \begin{bmatrix} \overline{r} + \tilde{r} \\ \overline{\theta} + \tilde{\theta} \end{bmatrix}$$

Thus:
$$\overline{y}_1 = \overline{r} \cos \overline{\theta}$$
 \therefore $\overline{y}_1 = 0$

2nd component:

 $\overline{y}_2 = E[r\sin\theta]$ $= \mathbf{E}\left[\left(\overline{r} + \widetilde{r}\right)\sin\left(\overline{\theta} + \widetilde{\theta}\right)\right]$ $= \mathbf{E} \left[\left(\overline{r} + \widetilde{r} \right) \left(\sin \overline{\theta} \cos \widetilde{\theta} - \cos \overline{\theta} \sin \widetilde{\theta} \right) \right]$ $= \overline{r} \sin \overline{\theta} \mathbf{E} \left[\cos \widetilde{\theta} \right]$ $= \mathbf{E} \left[\cos \tilde{\theta} \right]$

Assuming:

$$\mathbf{x} = \begin{bmatrix} r \\ \theta \end{bmatrix} = \begin{bmatrix} \overline{r} + \tilde{r} \\ \overline{\theta} + \tilde{\theta} \end{bmatrix}$$

Thus:
$$\overline{y}_2 = E\left[\cos\tilde{\theta}\right]$$
 but $\overline{y}_2 = 1???$

If we assume uniform distribution:

$$\theta \sim U\left[-\theta_m, \theta_m\right]$$

It follows that



The first order approximation is incorrect!!



How to improve this estimate?

In other words, we seek better ways to propagate

$$\overline{\mathbf{x}} = \mathbf{E}[\mathbf{x}]$$
 and $\mathbf{P} = \operatorname{cov}[\mathbf{x}]$

through a nonlinear mapping

$$\mathbf{y} = \mathbf{h}(\mathbf{x})$$

in order to obtain

$$\overline{\mathbf{y}} = \mathbf{E}[\mathbf{y}] \text{ and } \mathbf{P}_{y} = \operatorname{cov}[\mathbf{y}]$$

Solution: Unscented transform. 1st step (2n samples or sigma points): $\mathbf{x}^{(i)} = \overline{\mathbf{x}} + \widetilde{\mathbf{x}}^{(i)}$, with $\widetilde{\mathbf{x}}^{(i)} = \pm \left(\sqrt{n\mathbf{P}}\right)_{i=0.96}^{T=0.99}$ • n is the size of x vector;

• i subscript: ith row of matrix.

2nd step (Mapping of samples):

 $\mathbf{y}^{(i)} = \mathbf{h}\left(\mathbf{x}^{(i)}\right)$



Solution: Unscented transform. 3rd step (Averaging of mappings):

$$\overline{\mathbf{y}} = \frac{1}{2n} \sum_{i=1}^{2n} \mathbf{y}^{(i)}$$

and
$$\mathbf{P}_{y} = \frac{1}{2n} \sum_{i=1}^{2n} \Delta \mathbf{y} \Delta \mathbf{y}^{T}$$

4 ~ 4

OBS:
$$\Delta \mathbf{y}^{(i)} = \mathbf{y}^{(i)} - \overline{\mathbf{y}}$$

Unscented transform:

- 3rd order accurate;
- Nonintrusive;
- No extra matrices required;
- Sq. root of **P**: Cholesky decomp.;
- Deterministic sampling (2n samples)





The Unscented KF – Equations

Unscented KF (UKF): Unscented transform (UT) + Kalman Filter:

• 1 UT at prediction stage + 1 UT at update stage.

Prediction:

• Sampling of Sigma Points

$$\hat{\mathbf{x}}_{n-1}^{(i)} = \hat{\mathbf{x}}_{n-1}^{+} \pm \left(\sqrt{n\mathbf{P}_{k-1}^{+}}\right)_{i}^{T}$$
 and $\hat{\mathbf{x}}_{n}^{(i)} = \mathbf{f}\left(\hat{\mathbf{x}}_{n-1}^{(i)}, \mathbf{u}_{n}\right)$

• Prior Mean and Covariance

$$\hat{\mathbf{x}}_{n}^{-} = \frac{1}{2n} \sum_{i=1}^{2n} \hat{\mathbf{x}}_{n-1}^{(i)} \text{ and } \mathbf{P}_{n}^{-} = \frac{1}{2n} \sum_{i=1}^{2n} \Delta \hat{\mathbf{x}}_{n}^{(i)} \Delta \hat{\mathbf{x}}_{n}^{(i),T} + \mathbf{Q}_{n}$$

The Unscented KF – Equations

Update:

• Sampling of sigma points

$$\hat{\mathbf{x}}_{n}^{(i)} = \hat{\mathbf{x}}_{n}^{-} \pm \left(\sqrt{n\mathbf{P}_{k}^{-}}\right)_{i}^{T} \text{ and } \hat{\mathbf{y}}_{n}^{(i)} = \mathbf{h}\left(\hat{\mathbf{x}}_{n}^{(i)}\right)$$

1

Calculation of Intermediate Covariance Matrices

$$\hat{\mathbf{y}}_{n} = \frac{1}{2n} \sum_{i=1}^{2n} \hat{\mathbf{y}}_{n}^{(i)}$$
$$\mathbf{P}_{y} = \frac{1}{2n} \sum_{i=1}^{2n} \Delta \hat{\mathbf{y}}_{n}^{(i)} \Delta \hat{\mathbf{y}}_{n}^{(i),T} + \mathbf{R}_{n} \text{ and } \mathbf{P}_{xy} = \frac{1}{2n} \sum_{i=1}^{2n} \Delta \hat{\mathbf{x}}_{n}^{(i)} \Delta \hat{\mathbf{y}}_{n}^{(i),T}$$

2n

The Unscented KF – Equations

Update:

• Posterior Mean and Covariance:

$$\mathbf{K}_n = \mathbf{P}_{xy} \mathbf{P}_y^{-1}$$

$$\hat{\mathbf{x}}_n^+ = \hat{\mathbf{x}}_n^- + \mathbf{K}_n [\mathbf{y}_n - \hat{\mathbf{y}}_n]$$
 and $\mathbf{P}_n^+ = \mathbf{P}_n^- - \mathbf{K}_n \mathbf{P}_y \mathbf{K}_n^T$

Example V



Example V



Example VI – Introduction

Heat Flux Identification problem – revisited:

- Same model as example III;
- Estimation of heat flux at z=0;
- Temperature measurements at z=c;
- 3D, Nonlinear;
- Heat flux application from t=0;
- Use Complete Model for Inverse Problem solution;



Example VI – Results

Hot Spot #2: Temperature and heat flux estimates



Example VI – Results

Hot Spot #2: Temperature estimates and residuals



Example VI – Results

Computational time:

- Physical time: 2.0 s;
- Kalman filter: N/A;
- <u>UKF: 115200 s (approx. 32 h);</u>

UKF and use of the Complete Model:

- Inverse Analysis is made possible;
- Computational time is excessively large.

Overall Conclusions

Kalman filtering techniques

- Wide range of applications for sequential estimation;
- Much more variations available;

Classical Kalman Filter:

- Analytical solution for linear and Gaussian problems;
- Stable, robust, easy to program;

Extended and Unscented Kalman Filter:

- Reliable algorithms for nonlinear and/or non Gaussian problems;
- Increasing degrees of accuracy and complexity.

Overall Conclusions







(b) The above figure depicts the fact that the Kalman filter is optimal for linear Gaussian systems. Going from a Kalman filter to a UKF to a particle filter will increase computational effort but will not improve estimation accuracy.

SIMON D., Optimal State Estimation, John Wiley & Sons, 2006.